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INNOVATION AND THE EXPORT-PRODUCTIVITY LINK

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INNOVATION AND THE EXPORT-PRODUCTIVITY LINK

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Abstract

In this paper, we explore the relationship between innovation activity, productivity, and exports using a panel of Spanish manufacturing firms for 1990-1998. Our results – based on non-parametric tests – suggest that firm innovation status is critical in explaining the positive export-productivity association documented in prior research. For the sample of small innovating firms, we find no significant differences in productivity levels between exporters and non-exporters. Product innovation in particular seems to explain this positive association between exports and productivity. For small non-innovating firms with low and medium productivity levels, however, exporting firms continue to exhibit higher productivity than non-exporting firms.

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1. Introduction

The dynamic processes of firm formation, growth, prospering, and failure generate a considerable degree of heterogeneity in firm performance, not only across industries, but more interestingly, also within industries (Bartelsman and Doms, 2000).

Most of the theoretical models of industry dynamics assume that firms are born with an inherent ability, their productivity. Efficient firms survive and grow in the market, while inefficient firms, with productivity below a certain threshold, decline and fail (Jovanovic, 1982; Hopenhayn, 1992). These models, however, assume that the productivity distribution across firms is exogenous to firms, thus relating firm survival to luck-of-draw. Firms with low productivity exit, while “lucky” firms with high productivity survive and continue growing. Little room is left for firm decisions, except for the decision on exiting, which is endogenized. The model of Pakes and Ericson (1995) improves on these models by introducing investment decisions that can potentially enhance survival chances.

While theoretically such heterogeneity and dynamics is difficult to handle, empirically it provides a wealth of interesting observations. Nevertheless, we know very little about the connection between individual firm decisions and their dynamic consequences. One of the basic empirical facts related to productivity is a strong positive association between productivity and export activity at the firm level. Most studies explain this pattern by the self-selection of more efficient firms into the export market (Clerides, Lach, et al., 1998; Bernard and Jensen, 1999; Delgado, Farinas, et al., 2002; Fafchamps, El Hamine, et al. 2002), confirming the sunk cost hypothesis that only those firms who are efficient enough to bear the entry costs and intense competition of the export market will start exporting. This suggests that a closer examination of prior firm decisions might be needed to understand this important selection.

In this paper, we take the first step towards explaining the observed productivity–exports link. We argue that a potential underlying mechanism for the selection of more productive firms into exporting is related to firms’ innovation. Successful innovation enhances the firm’s productivity, leading to the selection of the more productive firms into the export markets. Yet,

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anything affecting its productivity could drive the firm into exports. Recent productivity literature, however, has found evidence that suggests that firm-specific demand variations, rather than technical efficiency, are the dominant factor in determining firm survival and positively influencing productivity (Foster et al., 2005). This suggests that product innovation related to positive demand shocks rather than innovation in processes related to production efficiency could be responsible for the increase in productivity and, consequently, entry into exporting. Consistent with this argument, product innovation has been found to play a very important role in explaining the firm's export decision (Basile, 2001; see related paper Cassiman and Martínez-Ros, 2007), showing that innovation-active firms are significantly more likely to become exporters than non-innovators. Thus, accounting for innovation may be critical in explaining the strong positive correlation between exporting and productivity in the existing research. We therefore argue that the observed productivity-export link may be partly explained by the firm's innovation status.

We examine the relationship between productivity and exports using a panel of Spanish manufacturing firms. We investigate the export-productivity link of the firms that engage in innovation activities and compare the results to those obtained for the non-innovating firms using non-parametric tests. As our findings indicate, innovating firms show no significant difference in productivity levels for exporting and non-exporting groups, suggesting that firm innovation strategy is a very important factor in explaining the exports-productivity association. Our findings have important policy implications. If innovation activity is a source of productivity growth, then policies aimed at promoting innovation, and product innovation in particular, might be more effective than direct export promotions, at least for firms "at risk" for innovating.

The paper proceeds as follows. In Section 2 we discuss the related literature. Sections 3 and 4 describe the data and methodology used in this study. Section 5 presents the results of the empirical analysis. A discussion section concludes.

2. Related Literature

In the empirical international trade literature, the positive association between exports and firm productivity has been well-documented. At least two explanations for the observed exports-productivity link have been suggested. On one hand, the positive association between exporting and productivity is explained through a selection mechanism. Sunk start-up costs associated with becoming an exporter, lead to the self-selection of more productive firms into exporting (Aw, Chen, and Roberts, 1997; Roberts and Tybout, 1997; Bernard and Jensen, 1999, 2004; Clerides, Lach and Tybout, 1998; Delgado, Farinas, et al., 2002; Fafchamps, El Hamine, et al., 2002). The hysteresis in exporting serves as evidence for the sunk entry costs in the export market. On the other hand, there is the possibility of learning-by-exporting – exporters may learn from their foreign contacts, adopting new production technologies and increasing productivity (Aw, Chung and Roberts, 2000; Delgado, Farinas, et al. 2002).

With both mechanisms being plausible, empirical evidence is rather unanimous in supporting the selection hypothesis behind the exports-productivity link (Roberts and Tybout, 1997; Clerides, Lach and Tybout, 1998; Bernard and Jensen, 1999; Delgado, Farinas, et al. 2002; Fafchamps, El Hamine, et al., 2002). The general finding is that exporting firms have higher

productivity than non-exporters before taking up exports and no significant productivity advantages are observed between continuous exporters and non-exporting firms over time.

Such heterogeneity in productivity raises an important question about the sources of these exporting firms' high productivity. How do firms obtain higher productivity levels that allow them to easily enter the export markets? International trade literature, following the work on industry dynamics (Jovanovic, 1982; Hopenhayn, 1992), has attempted to incorporate firm heterogeneity in international trade modeling. Recent theoretical work by Melitz (2003) and Bernard et al. (2003) formulates the theories that reflect the empirical regularities observed in export behavior and productivity. In these theories, a firm's initial productivity level is determined by a random draw from a certain distribution function. The model by Melitz (2003) assumes sunk entry costs in the export market, while Bernard et al. (2003) assume Bertrand competition among producers, which only allows the most productive firms to incur trading costs associated with exports. Thus, these theories demonstrate the selection mechanism of more productive firms into the export market. The models, however, do not explain why these firms are more productive and self-select into exporting, that is, the theories are not causal theories between firm decisions and a decision to export.

One important source of productivity differences seems to be related to R&D and innovation activities. A number of empirical studies have documented the positive and significant effect of R&D and innovation on firm productivity and productivity growth. Crepon et al. (1998), estimating a structural model that links productivity, innovation output and innovation inputs, find that firm productivity correlates positively with higher innovation output. In line with their result, Jefferson et al. (2002) for Chinese firms show that new product sales are positively associated with productivity. Using the panel of Spanish firms, Huergo and Jaumandreu (2004) find that process innovation is an important determinant of productivity growth at the firm level. Investigating the relationship between innovation and productivity in four European countries, Griffith et al. (2005) find results, consistent with the previous studies, that both product and process innovations have a positive significant effect on firm-level productivity in three out of the four countries.

At the same time, R&D and innovation activities seem to play a very important role in explaining firm's decision to export and export volumes. In particular, recent studies find that innovation is a very important driver of the export decision. For a sample of Italian manufacturing firms, Basile (2001) shows that firms introducing product and/or process innovations either through R&D or through investments in new capital are more likely to export. Bernard and Jensen (2004) find that changing primary SIC code – which could indicate the introduction of new products – significantly increases the probability of entering export markets. In a related paper, Cassiman and Martínez-Ros (2007) find a strong positive effect of product innovation on a firm's decision to export.

Taken together, prior empirical findings suggest that innovation activity may be responsible for both the productivity enhancement and export orientation of a firm and explain the correlation between exports and productivity. A number of studies provide empirical results going in the direction of our argument. Aw and Batra (1998), on a sample of Taiwanese firms, find that for the group of large, high-technology firms, exporters do not differ from non-exporters in their efficiency levels. However, in the group of small firms with no formal investments in technology, exporters are significantly closer to the production frontier than non-exporting firms. Using the sample of Spanish firms, Delgado et al. (2002) show that the export-productivity link varies depending on firm size. They observe no significant difference in

productivity levels between exporters and non-exporters for large firms. However, for small firms, exporters show significantly higher productivity levels than non-exporters. In a recent paper, Aw et al. (2005) find that for firms that do not invest in R&D, the exporters' productivity is significantly higher than that of non-exporting firms. Moreover, firms that export and invest in R&D are found to have higher productivity than those that only export. This evidence is used to argue that not only do more efficient firms select into the export market, but exports and R&D are important and complementary sources of productivity growth, with R&D activities facilitating the benefits from export markets. When coupled with the well-documented positive link between innovation and firm size (although possibly at a decreasing rate), these findings point to the importance of innovation as an explanatory variable driving the export-productivity link.

Therefore, connecting innovation, productivity and exports, we argue that accounting for innovation might take us some way in explaining the positive association between exports and productivity. Furthermore, in a recent paper, Foster et al. (2005) find that firm-specific demand shocks rather than production efficiency shocks explain differences in productivity, suggesting that product innovation rather than process innovation improves productivity, and, consequently, drives the decision to export.

3. Data

The data that are used in this study come from a survey of Spanish manufacturing firms started in 1990 with data collected annually up to 1998. The project was conducted by the Fundación Empresa Pública with financial support from the Spanish Ministry of Science and Technology. The information collected each year is consistent with the information in the previous years. The sample contains the population of firms with at least 200 employees and 4% of the population of firms with more than 10 and less than 200 employees. Firms that dropped out of the original sample are replaced every year by firm with the similar characteristics from the population. The survey contains detailed information for every year on exporting and innovation activities, reporting among other questions information on export volume and on product and process innovation carried out by a firm.¹

The initial sample includes 2188 firms in 1990 and 3195 firms in 1998. The dataset contains firms from 20 distinct industries and is representative of the Spanish manufacturing sector. Due to entry, exit, and missing values, the resulting sample includes 11,855 firm-year observations. The sample is an unbalanced panel, with a significant variation in the export and innovation behavior across firms as well as over time. Previous research [Delgado et al. (2002); Campa (2004); Huergo and Jaumandreu (2004); Cassiman and Martínez-Ros (2007), among others] has used the same data set as it is representative for the Spanish manufacturing sector.

4. Empirical Strategy and Methods

Our empirical strategy is as follows. We start by reproducing the results existing in the literature and identify the positive association between productivity and export status in our sample. Next, we compare the productivity levels of innovating versus non-innovating firms in

¹ The definitions of product and process innovation as they appear in the questionnaire are provided in Appendix.

order to show that innovation activity adds to firm productivity. Finally, we check whether the differences in the productivity of exporters and non-exporters persist when firm innovation status is taken into account.

Productivity Measure

To measure productivity we construct an index of total factor productivity for each firm, using a multilateral index developed by Caves et al. (1982) and extended by Good et al. (1997).² The TFP index is calculated as the logarithm of the firm's output less a cost-share weighted sum of the logarithms of the firm inputs. To make the comparison between any two firm-year observations possible, each firm's outputs and inputs are calculated as deviations from a reference firm. The reference firm is a hypothetical firm that varies across industries with input cost-based shares computed as an arithmetic mean of cost shares over all observations, and outputs and inputs computed as the geometric mean of outputs and inputs over all observations. Moreover, since the sampling proportions in our data are different for small (≤ 200 employees) and large firms (> 200 employees), the reference firm also varies across size groups. Thus, each firm's output, inputs and productivity for each year are measured relative to this hypothetical firm in the same size group (small or large) and industry. For more detail on the computation of the TFP index, see Appendix.

Methods

We start with the graphical description of the TFP distributions of exporting versus non-exporting, and innovating versus non-innovating firms across 1991-1998. In our analysis, we focus only on the small firms (≤ 200 employees), since the number of large firms in our sample is not sufficient for the test to be statistically conclusive.

Next, we conduct a number of tests to document the expected effects formally. We begin by comparing means and variances of the TFP level distributions of exporters and non-exporters. Then we compare the productivity distributions themselves across these two subsamples of firms. To test the differences in the TFP level distributions of exporters versus non-exporters, we employ a Kolmogorov-Smirnov equality-of-distributions test, used recently in Delgado et al. (2002). This non-parametric test rejects the null hypothesis of samples coming from the same populations if there is a point for which the cumulative empirical distributions of two independent samples are significantly different. The testing procedure is based on the concept of first-order stochastic dominance. Let F and G be cumulative distribution functions of TFP for two subsamples to be compared (in our case, e.g. exporters versus non-exporters). First-order stochastic dominance of F relative to G is defined as: $F(z)-G(z)\leq 0$ uniformly for any z from R , with strict inequality for at least one z . In order to show that F stochastically dominates G we need to conduct the following tests:

-two-sided test : $H_0: F(z)-G(z)=0$ for all z from R versus $H_a: F(z)-G(z)\neq 0$ for some z ;

-one-sided test: $H_0: F(z)-G(z) \leq 0$ for all z from R versus $H_a: F(z)-G(z)>0$ for some z .

² Our method of calculating the TFP index is similar to that performed in Aw et al. (2000) and Delgado et al. (2002).

The two-sided test checks the hypothesis on the equality of the distributions F and G. The distributions F and G are not significantly different if we cannot reject H_0 for the two-sided test. The one-sided test allows determining whether one distribution dominates the other. Not being able to reject H_0 for the one-sided test will mean that F is equal or to the right of the distribution G.

Thus, in order to show that F stochastically dominates G we have to demonstrate that the null hypothesis H_0 for the *two-sided test can be rejected*, while H_0 for the *one-sided test cannot be rejected*. This will be consistent with F being to the right of G. In our case, it will imply that the TFP level distribution of exporters stochastically dominates the distribution of TFP for non-exporters. We conduct the Kolmogorov-Smirnov test for each time period t , $t=1991, \dots, 1998$.

Next, we compare the productivity levels of innovators and non-innovators using the same battery of tests. We repeat the same tests for the exporting and non-exporting groups accounting for firms' innovation strategy. Finally, we run quantile regressions of productivity levels on export variable and several controls in order to investigate differences in the TFP levels of exporters and non-exporters in more detail.

We define exporters as firms exporting in the current year. Non-exporters are those firms that did not perform exports in the current year. Innovation activity is measured in several ways. We distinguish between innovating in product and in process, using two dummies that indicate whether a firm carried out a product or a process innovation. Next, we employ a dummy variable that indicates whether a firm has performed any innovation activity (either product or process). Finally, we use a measure for the innovation input – whether a firm invested in R&D. We measure our innovation variables with a one-year lag, since innovation is unlikely to drive the productivity improvements in the same year.³

5. Results

We start with the graphical representation of the cumulative distribution functions of TFP levels for the different groups of firms, looking at the productivity distributions of 1) exporting versus non-exporting firms; 2) innovating versus non-innovating firms; and 3) exporters versus non-exporters for the innovating and non-innovating groups.

Figures 1-5 present the results for these subsamples of firms. The distribution of performers (exporters or innovators) lies to the right of the distribution of non-performers, which suggests first-order stochastic dominance. The exception is the process innovation case, for which TFP level distributions of innovating and non-innovating firms seem to coincide. Figures 6-9 compare the productivity distributions of exporters and non-exporters in the groups of innovators and non-innovators. For the non-innovating firms, the TFP distribution of exporters is clearly to the right of that of non-exporters, which points to stochastic dominance. In the group of innovating firms the difference between TFP distributions is not so evident, especially for the product innovation case. Overall, the visual comparison of the TFP level distributions shows that the productivity distribution of exporters dominates that of non-exporters, which

³ See Appendix for the description of the variables.

also holds in the group of non-innovators. For innovating firms, however, the difference is less pronounced, hinting at the existence of the hypothesized effect of innovation activity on productivity and export decision. In the following, we perform a formal comparison of TFP level distributions.

5.1. Exporters Versus Non-Exporters

First, we formally document the existence of the positive association between export orientation and firm productivity in line with prior research.

Table 1 lists the results for tests on means and variances and the Kolmogorov-Smirnov test for exporters and non-exporters.

Exporting Versus Non-Exporting Firms

Table 1

Difference in TFP Level Distributions Between Exporters and Non-Exporters

	Number of firms		Difference in means	Mean (non-exporters) > mean (exporters)		Variance (non-exporters) > variance (exporters)		K-S test for equality of distributions*		Difference in favor of exporters	
	exp=1	exp=0		t-statistic	P-value	F-statistic	P-value	D	P-value	D	P-value
1991	295	591	-0.011	-0.694	0.244	1.433	0.999	0.067	0.303	-0.032	0.667
1992	378	629	-0.030	-2.067	0.019	1.320	0.998	0.104	0.009	-0.025	0.752
1993	377	604	-0.051	-3.809	0.000	1.425	0.999	0.120	0.002	-0.004	0.993
1994	439	586	-0.044	-3.371	0.000	0.974	0.384	0.124	0.001	-0.008	0.964
1995	449	504	-0.059	-4.271	0.000	1.046	0.690	0.132	0.000	-0.001	1.000
1996	483	532	-0.057	-4.176	0.000	1.074	0.787	0.174	0.000	-0.007	0.975
1997	616	585	-0.073	-5.809	0.000	1.188	0.982	0.169	0.000	0.000	1.000
1998	573	528	-0.067	-5.076	0.000	1.490	1.000	0.142	0.000	-0.002	0.997

* Here and further on, we test the following two hypotheses:

Ho(1): $F(\text{TFP group1})=F(\text{TFP group2})$ – the test for the equality of distributions;

Ho(2): $F(\text{TFP group1})<F(\text{TFP group2})$ – the test for the differences in TFP levels favorable to group1, where group1 and group2 are the groups of exporting and non-exporting (or innovating and non-innovating) firms, respectively.

The third and the fourth columns report the difference in means and the test statistic on the null hypothesis that the mean TFP level of non-exporting firms is significantly higher than the mean TFP level of exporting firms. The comparison of average TFP levels indicates that exporting firms have higher levels of TFP than non-exporters, with the difference being statistically significant.

The significance of the results, however, varies across years. During 1991-1998, Spain underwent the entire business cycle, with a slowdown in the economy starting in 1991, a sharp recession in 1993 – beginning of 1994, and a recovery in 1995-1998. The unfavorable economic conditions in the early nineties might be responsible for the non-significance of test

statistics when comparing the TFP distributions of exporters/non-exporters and innovators/non-innovators.⁴

Column five presents the test statistic for the hypothesis of greater variability of TFP level for non-exporting firms than for exporting ones. In most cases, we cannot reject the hypothesis that the variance of TFP levels of non-exporting firms is larger than that of exporters

Columns six and seven report the results for the Kolmogorov-Smirnov test – the statistic for the two-sided test on the equality of distributions and the one-sided test results. We can reject the null hypothesis of equality of distributions for exporters and non-exporters at 1% significance level. In the one-sided test, the null hypothesis states that the TFP distribution of exporters stochastically dominates the TFP distribution of non-exporters. As the results in column 7 show, the null hypothesis for the one-sided test, i.e. that the TFP level differences are in favor of exporters, cannot be rejected.

Therefore, the results in Table 1 confirm the findings of the prior studies for our sample, showing that firm export status is indeed associated with higher productivity levels. Exporting firms not only show higher levels of TFP, but the distributions of TFP for exporters and non-exporters are significantly different, with the exporters' TFP distribution stochastically dominating the non-exporters' TFP distribution.

5.2. Innovators Versus Non-Innovators

We further explore the differences in productivity levels between innovating and non-innovating firms. Table 2A-D lists the results for product and process innovation variables, as well as for innovation and R&D dummies.

Innovating Versus Non-Innovating Firms

Table 2A

Difference in TFP Level Distributions Between Firms with Product Innovation and Firms with No Innovation

	Number of firms		Difference in means	Mean (no innovation) > mean (prod. innovation)		Variance (no innovation) > variance (prod. innovation)		K-S test for equality of distributions		Difference in favor of firms with prod. innovation	
	Prod. innovation	No innovation		t-statistic	P-value	F-statistic	P-value	D	P-value	D	P-value
1991	60	689	-0.032	-1.099	0.136	1.128	0.713	0.127	0.284	-0.050	0.757
1992	90	588	-0.016	-0.603	0.274	1.437	0.983	0.110	0.252	-0.029	0.879
1993	98	558	-0.019	-0.812	0.209	1.331	0.959	0.119	0.155	-0.050	0.658
1994	109	593	-0.046	-2.095	0.018	1.240	0.916	0.155	0.017	-0.031	0.835
1995	100	587	-0.055	-2.363	0.009	1.116	0.747	0.131	0.082	-0.012	0.976
1996	109	558	-0.050	-2.177	0.015	1.164	0.834	0.146	0.031	-0.007	0.992
1997	103	605	-0.044	-1.858	0.032	1.167	0.832	0.133	0.069	-0.009	0.986
1998	121	657	-0.036	-1.702	0.044	1.191	0.882	0.110	0.139	-0.021	0.915

⁴ In general, we observe non-significant results for the early 90's, while testing the differences in the TFP distributions for both export and innovation variables. However, we still observe lower productivity levels in the groups of non-performers.

Table 2B

Difference in TFP Level Distributions Between Firms with Process Innovation and Firms with No Innovation

	Number of firms		Difference in means	Mean (no innovation)> mean (proc. innovation)		Variance (no innovation)> variance (proc. innovation)		K-S test for equality of distributions		Difference in favor of firms with proc. innovation	
	Proc. innovation	No innovation		t-statistic	P-value	F-statistic	P-value	D	P-value	D	P-value
1991	41	689	0.028	0.815	0.792	1.086	0.612	0.114	0.626	-0.114	0.365
1992	165	588	-0.002	-0.109	0.457	1.207	0.927	0.050	0.887	-0.050	0.528
1993	143	558	-0.004	0.227	0.589	1.328	0.979	0.053	0.882	-0.053	0.524
1994	159	593	0.010	0.543	0.706	1.107	0.779	0.088	0.255	-0.088	0.146
1995	153	587	-0.014	-0.717	0.237	0.862	0.116	0.126	0.034	-0.033	0.765
1996	151	558	-0.019	-0.926	0.177	1.105	0.768	0.102	0.143	-0.033	0.775
1997	151	605	-0.021	-1.053	0.146	1.382	0.992	0.094	0.206	-0.015	0.948
1998	205	657	-0.009	-0.503	0.308	0.741	0.003	0.063	0.522	-0.032	0.730

Table 2C

Difference in TFP Level Distributions Between Firms with Innovation (Either Product or Process) and Firms with No Innovation

	Number of firms		Difference in means	Mean (no innovation)> mean (innovation)		Variance (no innovation)> variance (innovation)		K-S test for equality of distributions		Difference in favor of firms with innovation	
	Innovation	No innovation		t-statistic	P-value	F-statistic	P-value	D	P-value	D	P-value
1991	160	689	-0.028	0.814	0.066	1.262	0.963	0.095	0.161	-0.031	0.779
1992	372	588	-0.015	-1.083	0.139	1.306	0.997	0.071	0.172	-0.016	0.889
1993	348	558	-0.005	-0.362	0.358	1.259	0.990	0.055	0.494	-0.043	0.439
1994	372	594	-1.444	-0.020	0.074	1.157	0.938	0.057	0.402	-0.019	0.844
1995	361	588	-0.029	-2.011	0.022	1.012	0.549	0.108	0.008	-0.016	0.883
1996	351	558	-0.024	-1.667	0.047	1.103	0.843	0.089	0.055	-0.021	0.826
1997	342	604	-0.031	-2.139	0.016	1.252	0.989	0.103	0.016	-0.006	0.981
1998	443	657	-0.026	-2.020	0.021	0.961	0.325	0.085	0.037	-0.008	0.967

Table 2D

Difference in TFP Level Distributions Between Firms with R&D and Firms with No R&D

	Number of firms		Difference in means	Mean (R&D=0)> mean (R&D=1)		Variance (R&D=0)> variance (R&D=1)		K-S test for equality of distributions		Difference in favor of firms with R&D	
	R&D =1	R&D =0		t-statistic	P-value	F-statistic	P-value	D	P-value	D	P-value
1991	144	705	-0.021	-1.099	0.136	1.047	0.627	0.130	0.027	-0.015	0.951
1992	198	762	-0.029	-1.615	0.053	1.417	0.998	0.115	0.024	-0.020	0.885
1993	185	721	-0.023	-1.371	0.085	1.452	0.999	0.092	0.140	-0.039	0.633
1994	206	759	-0.028	-1.715	0.043	1.012	0.533	0.076	0.273	-0.006	0.989
1995	204	745	-0.035	-2.059	0.019	1.303	0.989	0.116	0.021	-0.017	0.915
1996	181	729	-0.045	-2.469	0.007	1.304	0.985	0.140	0.005	-0.029	0.788
1997	197	749	-0.039	-2.272	0.012	1.217	0.953	0.118	0.020	-0.019	0.894
1998	232	868	-0.036	-2.271	0.012	1.295	0.992	0.110	0.020	-0.018	0.892

On average, innovation-active firms show higher productivity levels than non-innovating ones. Firms engaged in product innovation have a higher average TFP level. Contrary to other findings in the literature, process innovation does not seem to have a similar significant effect on productivity. Variability of TFP levels for innovating and non-innovating groups does not reveal any recognizable pattern, especially in the case of process innovation. The results for the Kolmogorov-Smirnov two-sided and one-sided tests show that the null hypothesis of equal distributions can be rejected for product innovators. Moreover, lower productivity levels are observed in the group of non-innovators. Process innovation does not seem to have a differential effect on productivity levels of innovators relative to non-innovators.

In Table 2C-D, the results for the innovation and R&D variables are presented. Again, on average, innovating firms show higher TFP levels compared to firms in the non-innovating group. The comparison of average TFP levels indicates that firms engaged in innovation have higher total factor productivity levels. These firms also show lower variability in TFP. Kolmogorov-Smirnov one-sided and two-sided tests show that the null hypothesis of equal distributions can be rejected and lower productivity levels are observed in the group of firms with no innovation.

5.3. Exporters Versus Non-Exporters Conditional on Innovation

Finally, we conduct the tests for the subsamples of exporters and non-exporters accounting for firm innovation status. The results of the tests are listed in Tables 3A-E.

Exporters Versus Non-Exporters Conditional on Firm Innovation Status

Table 3A

Differences in TFP Level Distributions Between Exporters and Non-Exporters Conditional on Firm Innovation Status. Product Innovation

Product innovation =1											
	Number of firms		Difference in means	Mean (non-exporters)> mean (exporters)		Variance (exporters)> variance (non-exporters)		K-S test for equality of distributions		Difference in favor of exporters	
	exp=1	exp=0		t-statistic	P-value	F-statistic	P-value	D	P-value	D	P-value
1991	27	33	0.096	1.863	0.966	3.105	0.997	0.319	0.060	-0.319	0.048
1992	57	33	0.011	0.265	0.604	0.915	0.401	0.148	0.670	-0.148	0.399
1993	58	40	-0.024	-0.631	0.264	1.437	0.895	0.132	0.740	-0.132	0.439
1994	67	42	-0.028	-0.732	0.232	0.952	0.440	0.114	0.845	-0.064	0.808
1995	68	32	-0.036	-0.822	0.206	1.260	0.787	0.159	0.546	-0.093	0.682
1996	74	35	-0.063	-1.493	0.069	1.253	0.791	0.234	0.101	-0.057	0.854
1997	76	27	0.003	0.068	0.527	1.394	0.865	0.126	0.864	-0.099	0.672
1998	85	36	-0.100	-2.623	0.004	0.585	0.039	0.285	0.020	-0.031	0.952

Table 3B

Differences in TFP Level Distributions Between Exporters and Non-Exporters Conditional on Firm Innovation Status. Process Innovation

Process innovation=1											
	Number of firms		Difference in means	Mean (non-exporters)> mean (exporters)		Variance (non-exporters)> variance (exporters)		K-S test for equality of distributions		Difference in favor of exporters	
	exp=1	exp=0		t-statistic	P-value	F-statistic	P-value	D	P-value	D	P-value
1991	16	25	0.030	0.457	0.675	1.651	0.842	0.240	0.517	-0.240	0.325
1992	65	100	-0.027	-0.801	0.212	0.734	0.083	0.103	0.734	-0.031	0.928
1993	86	57	-0.032	-1.037	0.150	0.917	0.355	0.118	0.651	-0.065	0.745
1994	73	86	-0.028	-0.863	0.194	0.958	0.424	0.108	0.687	-0.063	0.729
1995	73	79	-0.075	-1.988	0.024	1.117	0.682	0.182	0.120	-0.014	0.984
1996	69	81	-0.022	-0.642	0.260	0.748	0.110	0.096	0.842	-0.043	0.869
1997	80	71	-0.052	-1.691	0.046	1.271	0.850	0.163	0.212	-0.014	0.985
1998	112	93	-0.105	-3.074	0.001	1.512	0.981	0.244	0.003	0.000	1.000

Table 3C

Differences in TFP Level Distributions Between Exporters and Non-Exporters Conditional on Firm Innovation Status. Innovation (Product and/or Process)

Innovation=1											
	Number of firms		Difference in means	Mean (non-exporters)> mean (exporters)		Variance (non-exporters)> variance (exporters)		K-S test for equality of distributions		Difference in favor of exporters	
	exp=1	exp=0		t-statistic	P-value	F-statistic	P-value	D	P-value	D	P-value
1991	77	83	0.030	0.997	0.840	2.102	0.999	0.167	0.165	-0.167	0.106
1992	191	181	-0.030	-1.419	0.078	0.919	0.284	0.082	0.504	-0.008	0.988
1993	179	169	-0.032	-1.620	0.053	1.148	0.819	0.068	0.774	-0.006	0.992
1994	212	160	-0.032	-1.571	0.058	0.925	0.304	0.079	0.560	-0.014	0.964
1995	214	147	-0.059	-2.565	0.005	1.206	0.893	0.133	0.071	-0.006	0.993
1996	218	133	-0.037	-1.606	0.054	0.783	0.062	0.128	0.103	-0.013	0.969
1997	212	130	-0.022	-1.010	0.156	1.060	0.649	0.119	0.167	-0.015	0.963
1998	278	165	-0.086	-4.063	0.000	1.302	0.972	0.178	0.002	0.000	1.000

Table 3D

Differences in TFP Level Distributions Between Exporters and Non-Exporters Conditional on Firm Innovation Status. No-Innovation Case

No-innovation case											
	Number of firms		Difference in means	Mean (non-exporters)> mean (exporters)		Variance (non-exporters)> variance (exporters)		K-S test for equality of distributions		Difference in favor of exporters	
	exp=1	exp=0		t-statistic	P-value	F-statistic	P-value	D	P-value	D	P-value
1991	213	476	-0.007	-0.423	0.336	1.014	0.542	0.065	0.507	-0.020	0.886
1992	181	407	-0.024	-1.151	0.125	1.526	0.999	0.120	0.042	-0.065	0.345
1993	177	381	-0.067	-3.527	0.000	1.644	0.999	0.148	0.007	-0.030	0.801
1994	196	398	-0.048	-2.581	0.005	1.011	0.530	0.157	0.002	-0.007	0.984
1995	231	357	-0.056	-3.079	0.001	0.972	0.403	0.165	0.001	-0.003	0.997
1996	221	338	-0.073	-3.826	0.000	1.164	0.889	0.201	0.000	-0.011	0.967
1997	275	330	-0.076	-4.178	0.000	0.948	0.322	0.184	0.000	-0.003	0.997
1998	295	362	-0.046	-2.743	0.003	1.555	1.000	0.129	0.006	-0.008	0.978

Table 3E

Differences in TFP Level Distributions Between Exporters and Non-Exporters Conditional on Firm R&D Status

R&D =1											
	Number of firms		Difference in means	Mean (non-exporters)> mean (exporters)		Variance (non-exporters)> variance (exporters)		K-S test for equality of distributions		Difference in favor of exporters	
	exp=1	exp=0		t-statistic	P-value	F-statistic	P-value	D	P-value	D	P-value
1991	100	44	-0.001	-0.015	0.493	1.284	0.845	0.130	0.597	-0.060	0.797
1992	137	61	-0.006	-0.226	0.410	0.980	0.475	0.056	0.999	-0.048	0.819
1993	133	52	-0.0310	-1.085	0.139	1.702	0.991	0.178	0.140	-0.072	0.678
1994	143	63	-0.047	-1.512	0.065	0.986	0.486	0.125	0.431	-0.027	0.938
1995	159	45	-0.058	-1.787	0.037	1.494	0.961	0.176	0.173	-0.038	0.903
1996	138	43	-0.077	-2.291	0.011	0.584	0.023	0.281	0.006	-0.029	0.946
1997	147	50	-0.045	-1.384	0.083	1.143	0.731	0.153	0.281	-0.033	0.921
1998	174	58	-0.066	-2.279	0.011	0.996	0.508	0.2011	0.041	-0.005	0.997
R&D =0											
	Number of firms		Difference of means	Mean (non-exporters)> mean (exporters)		Variance (non-exporters)> variance (exporters)		K-S test for equality of distributions		Difference in favor of exporters	
	exp=1	exp=0		t-statistic	P-value	F-statistic	P-value	D	P-value	D	P-value
1991	190	515	0.004	0.230	0.590	1.175	0.904	0.033	0.997	-0.033	0.731
1992	235	527	-0.027	-1.551	0.060	1.279	0.984	0.090	0.118	-0.033	0.700
1993	223	498	-0.056	-3.344	0.000	1.303	0.988	0.110	0.037	-0.001	1.000
1994	265	495	-0.039	-2.459	0.007	1.019	0.565	0.128	0.005	-0.010	0.964
1995	286	459	-0.056	-3.411	0.000	0.923	0.224	0.135	0.002	-0.000	1.000
1996	302	427	-0.052	-3.138	0.000	1.037	0.632	0.149	0.001	-0.006	0.983
1997	340	409	-0.058	-3.631	0.000	0.965	0.367	0.156	0.000	-0.001	0.999
1998	399	469	-0.059	-3.968	0.000	1.451	0.999	0.140	0.000	-0.003	0.993

In line with our hypothesis, we find that innovating exporters do not differ significantly in their productivity levels from innovating non-exporters. Kolmogorov-Smirnov test statistics support these findings – we cannot reject the null hypothesis of equal distributions in the group of innovators, although exporting innovators still show higher productivity levels.

For the non-innovating group, however, the results are drastically different. We observe a significant difference in TFP levels for exporters and non-exporters. The tests indicate that exporting firms outperform non-exporting ones with respect to TFP showing significantly higher levels of productivity. Moreover, the two-sided Kolmogorov-Smirnov test statistic rejects the hypothesis of equality of distributions.

5.4. Quantile regression results

To get a more complete picture of the way productivity is linked to exports we use quantile regressions. More specifically, we regress the TFP variable on the export dummy and control variables including foreign capital ownership, high-tech sector and year dummies. We also test the association between the TFP level and innovation activity, regressing TFP on innovation variables.

The association between export/innovation status and productivity may vary at different points of the conditional distribution of productivity levels, and a quantile regression provides information on this variation. For each quantile, it can be shown whether the association between exports/innovation and productivity is positive, negative or insignificant, and how strong it is compared to other quantiles.

Table 4A-C presents the regression estimates of the export/innovation variables for OLS regression and five different quantiles of the TFP level distribution for 1991-1998.

Quantile Regression Results, 1991-1998⁵

Table 4A

Estimated Results of TFP Levels for the Export/Innovation Dummy Variables, 1991-1998

Dependent variable: TFP level						
Reported coefficient: Export (0/1)						
	OLS	Quantile regression				
Export (0/1)		5%	25%	50%	75%	95%
	0.044*** (0.008)	0.084*** (0.012)	0.052*** (0.005)	0.033*** (0.004)	0.018*** (0.005)	0.043*** (0.013)

Table 4B

Estimated Results of TFP Levels for the Innovation Variables, 1991-1998

Dependent variable: TFP level						
Reported coefficient: Product innovation (0/1)						
	OLS	Quantile regression				
Product innovation (0/1)		5%	25%	50%	75%	95%
	0.032*** (0.009)	0.057** (0.022)	0.045*** (0.009)	0.019*** (0.007)	0.011 (0.009)	0.032 (0.025)

Dependent variable: TFP level						
Reported coefficient: Process innovation (0/1)						
	OLS	Quantile regression				
Process innovation (0/1)		5%	25%	50%	75%	95%
	0.003 (0.007)	-0.005 (0.021)	0.010 (0.007)	0.000 (0.007)	0.006 (0.006)	-0.018 (0.019)

⁵ Standard errors are given in brackets. Foreign capital and high-tech industry dummy are included as covariates. Year fixed effects are included. *, **, *** are significantly different from zero at the 10%, 5% or 1% level, respectively.

Dependent variable: TFP level						
Reported coefficient: Innovation (0/1)						
	OLS	Quantile regression				
Process innovation (0/1)		5%	25%	50%	75%	95%
	0.018*** (0.005)	0.035** (0.015)	0.029*** (0.006)	0.017*** (0.005)	0.011** (0.004)	-0.004 (0.014)

Dependent variable: TFP level						
Reported coefficient: RD (0/1)						
	OLS	Quantile regression				
RD (0/1)		5%	25%	50%	75%	95%
	0.025*** (0.007)	0.050*** (0.013)	0.035*** (0.004)	0.028*** (0.004)	0.019*** (0.005)	-0.014 (0.012)

Table 4C

Estimated Results of TFP Levels for the Export Dummy Variable Conditional on Firm Innovation Status, 1991-1998

Product innovation=1						
Dependent variable: TFP level						
Reported coefficient: Export (0/1)						
	OLS	Quantile regression				
Export (0/1)		5%	25%	50%	75%	95%
	0.017 (0.018)	0.041 (0.049)	0.025 (0.016)	-0.002 (0.014)	0.000 (0.014)	0.043 (0.057)

Process innovation=1						
Dependent variable: TFP level						
Reported coefficient: Export (0/1)						
	OLS	Quantile regression				
Export (0/1)		5%	25%	50%	75%	95%
	0.044*** (0.017)	0.097*** (0.032)	0.034* (0.019)	0.035*** (0.013)	0.013 (0.018)	0.076*** (0.027)

Innovation=1						
Dependent variable: TFP level						
Reported coefficient: Export (0/1)						
	OLS	Quantile regression				
Export (0/1)		5%	25%	50%	75%	95%
	0.035*** (0.012)	0.084*** (0.016)	0.024*** (0.009)	0.026*** (0.008)	0.009 (0.010)	0.074*** (0.022)

No-innovation case						
Dependent variable: TFP level						
Reported coefficient: Export (0/1)						
	OLS	Quantile regression				
Export (0/1)		5%	25%	50%	75%	95%
	0.042*** (0.010)	0.070*** (0.018)	0.053*** (0.008)	0.035*** (0.006)	0.017** (0.007)	0.021 (0.019)

R&D=0						
Dependent variable: TFP level						
Reported coefficient: Export (0/1)						
	OLS	Quantile regression				
Export (0/1)		5%	25%	50%	75%	95%
	0.040*** (0.010)	0.073*** (0.015)	0.046*** (0.006)	0.030*** (0.006)	0.012** (0.005)	0.052*** (0.016)
R&D=1						
Dependent variable: TFP level						
Reported coefficient: Export (0/1)						
	OLS	Quantile regression				
Export (0/1)		5%	25%	50%	75%	95%
	0.038** (0.018)	0.111*** (0.036)	0.041*** (0.015)	0.027** (0.011)	0.014 (0.015)	0.029 (0.028)

We find that the significance and the magnitude of the export coefficient vary considerably as we move from the lower quantile (0.05) to the upper quantile (0.95) of the conditional productivity distribution. The association between productivity and exports seems to be strengthened in the lower tail and the center of the distribution but weakened towards the upper tail, suggesting that among the most productive firms, exporters and non-exporters tend to vary less in their productivity levels. With respect to the innovation variables, the results support our previous finding that product innovation is associated with the higher productivity, while process innovation comes out insignificant.

Accounting for firms' innovation status, in the group of non-innovators we find a positive and significant association between exports and the productivity level. The positive correlation between the export variable and productivity remains significant for the firms engaged in process innovation. For the firms performing product innovation, however, the export coefficient is insignificant along the entire productivity distribution, suggesting the importance of the effect of product innovation in explaining the export-productivity link. Thus, the results show that it is mainly product innovation that accounts for the differences in the productivity levels and consequently leads firms to export.

6. Discussion

In this paper, we examine the relationship between exports, productivity, and innovation at the firm level. Our findings highlight that the positive link between exports and productivity differs considerably depending on the firm's innovation strategy. We do not observe significant differences in productivity levels between exporters and non-exporters among firms that carry out product innovation. Once we take into account the innovation strategy, firm productivity comes out to be independent of whether or not a firm participates in exports.

However, the positive link between exports and productivity observed in prior research does exist for non-innovating firms, consistent with the learning-by-exporting effect emphasized by recent studies. The observed superior efficiency among exporting firms in the low and middle-productivity range may be related to their ability to get new technological information on the export markets and to the higher competition abroad. For the most productive firms, however, the positive association between exports and productivity is found to be weaker, suggesting

that for these firms the learning-by-exporting effect might not be strong enough to affect their initial productivity levels.

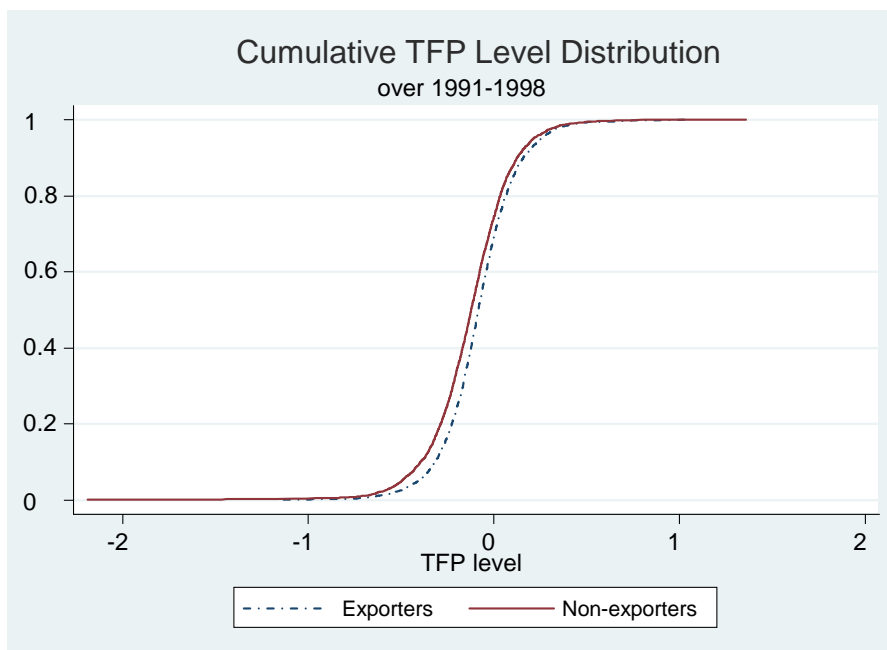
Our results suggest that innovation, and more specifically product innovation, allows firms to enter the export market. Successful product innovation enhances the firm's productivity leading to the selection of the more productive firm into the export markets. This finding appears to be especially relevant from a public policy perspective. If innovation activity is a source of productivity growth, then policies aimed at promoting innovation, and product innovation in particular, might be more effective than direct export promotions, at least for firms "at risk" for innovating.

There remain several issues to address. More specifically, we would like to look at the evolution of firms over time comparing different groups of firms, such as "always" exporting firms, non-exporting firms, firms entering in and exiting from export markets. We will focus on the subsample of non-innovating firms, thus isolating the effect of innovation, while testing for the selection versus learning hypothesis in the productivity-export link.

Exporting Versus Non-Exporting Firms

Figure 1

Cumulative Distribution of TFP Levels. Exporters Versus Non-Exporters, 1991-1998



Innovating Versus Non-Innovating Firms

Figure 2

Cumulative Distribution of TFP Levels. Firms with Product Innovation Versus Firms with No Innovation, 1991-1998

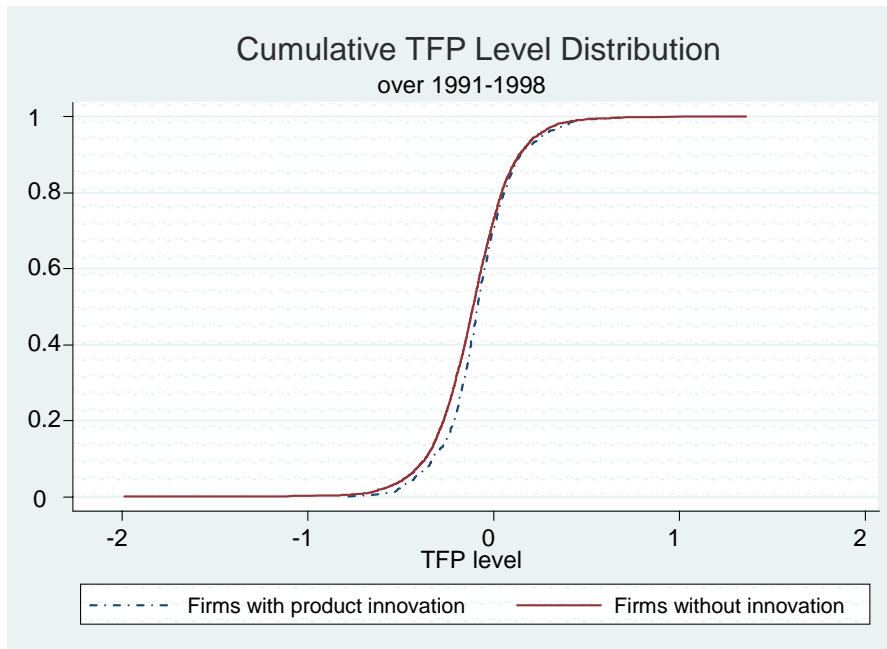


Figure 3

Cumulative Distribution of TFP Levels. Firms with Process Innovation Versus Firms with No Innovation, 1991-1998

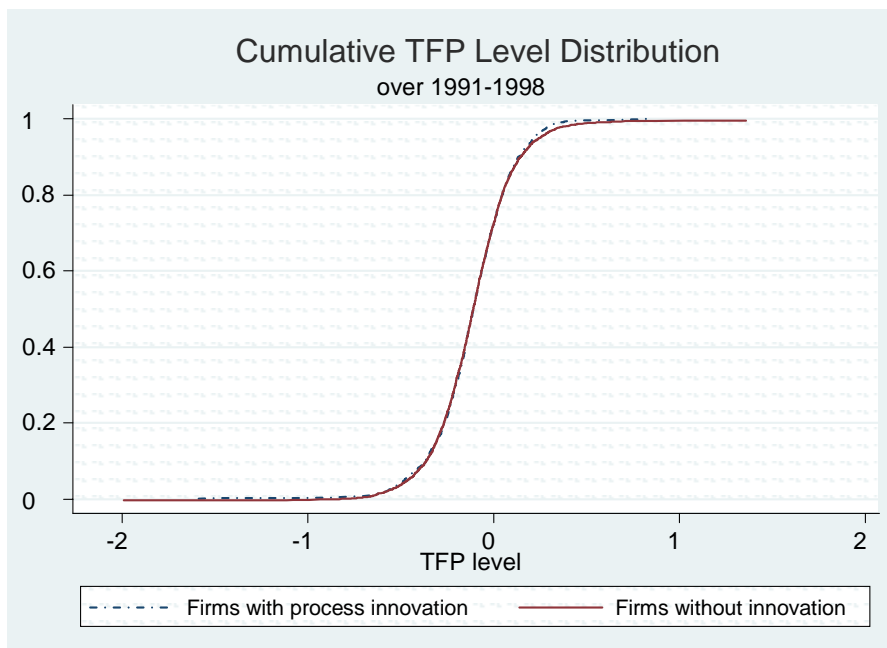


Figure 4

Cumulative Distribution of TFP Levels. Firms with Innovation Versus Firms with No Innovation, 1991-1998

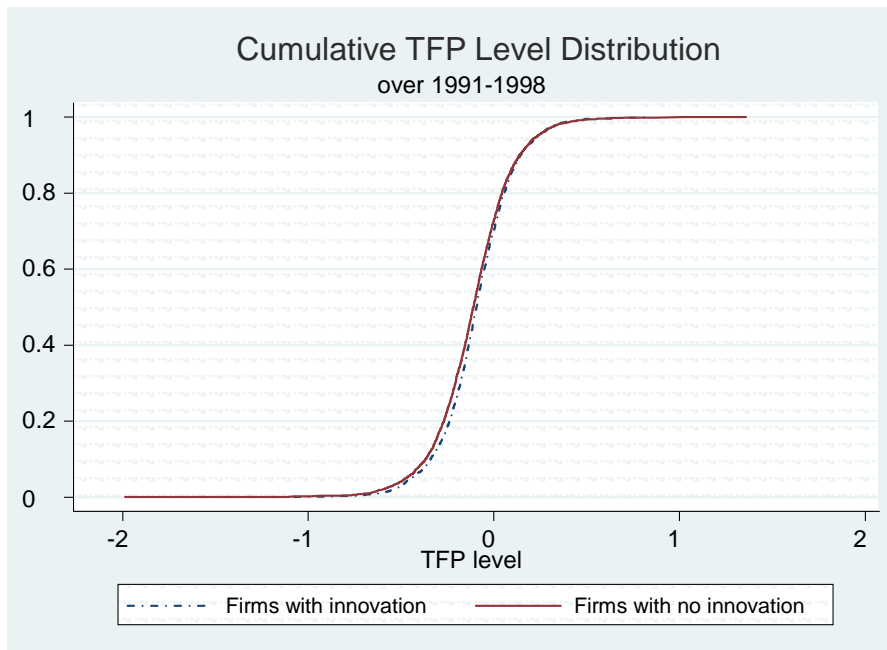
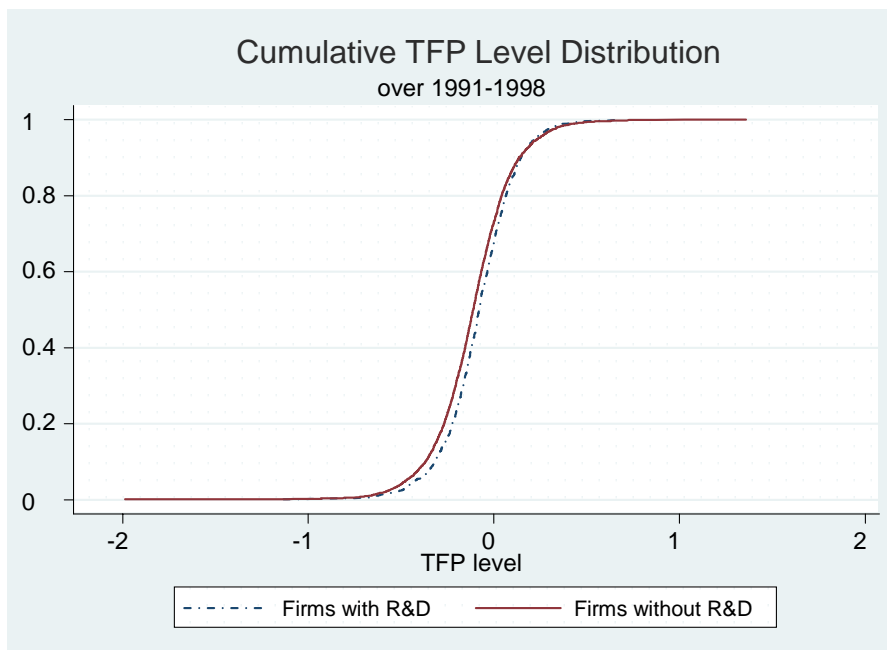


Figure 5

Cumulative Distribution of TFP Levels. Firms with R&D Versus Firms without R&D, 1991-1998

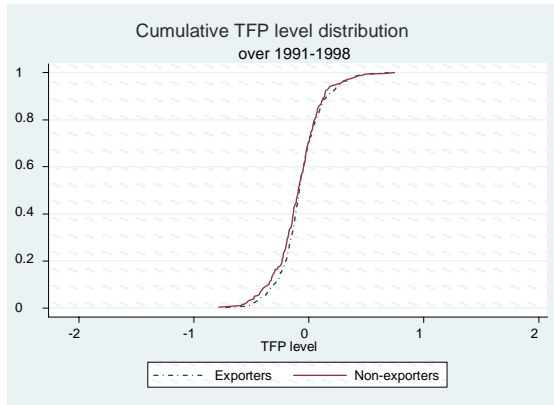


Exporting Versus Non-Exporting Firms Conditional on Innovation Status

Figure 6

Cumulative Distribution of TFP Levels. Exporters Versus Non-Exporters, for Firms with Product Innovation and without Innovation, 1991-1998

Product innovation=1



No-innovation case

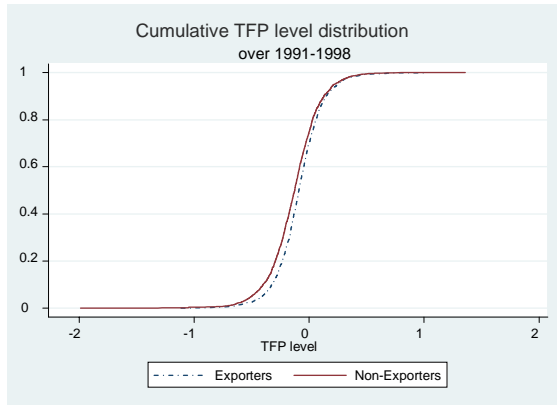
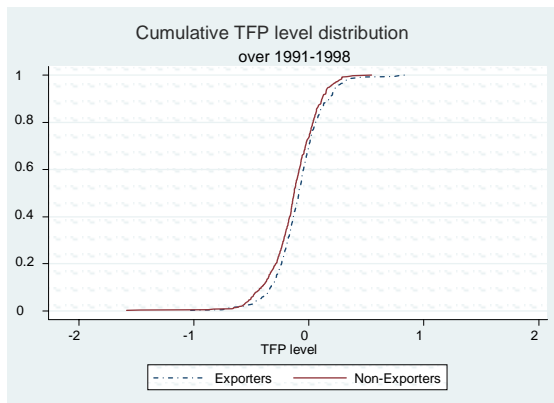


Figure 7

Cumulative Distribution of TFP Levels. Exporters Versus Non-Exporters, for Firms with Process Innovation and without Innovation, 1991-1998

Process innovation =1



No-innovation case

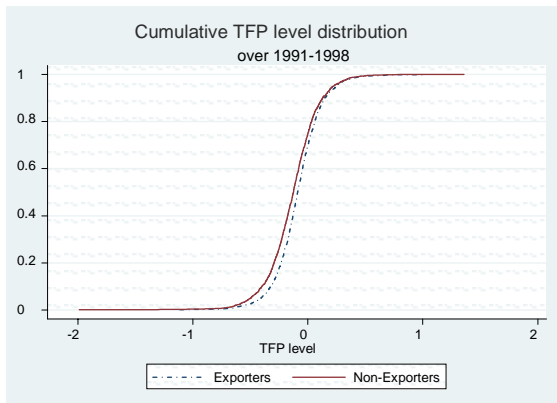
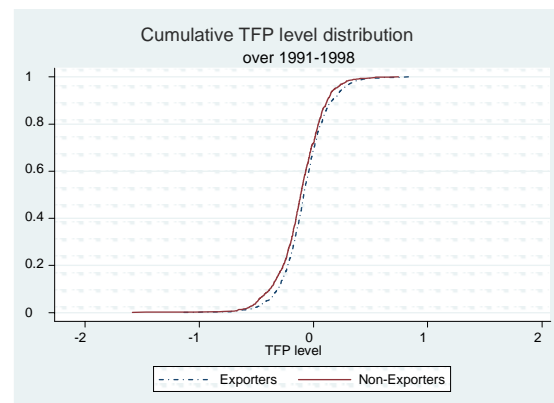


Figure 8

Cumulative Distribution of TFP Levels. Exporters Versus Non-Exporters, for Firms with Innovation and without Innovation, 1991-1998

Innovation =1



No-innovation case

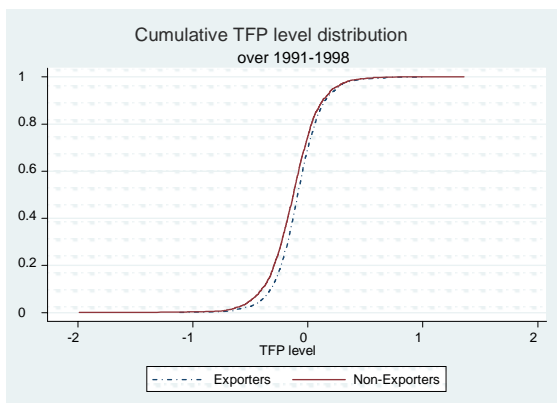
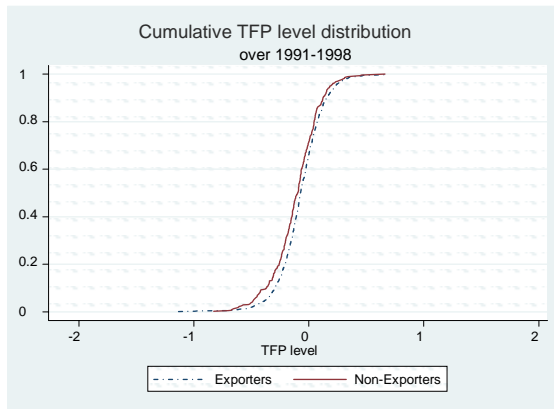


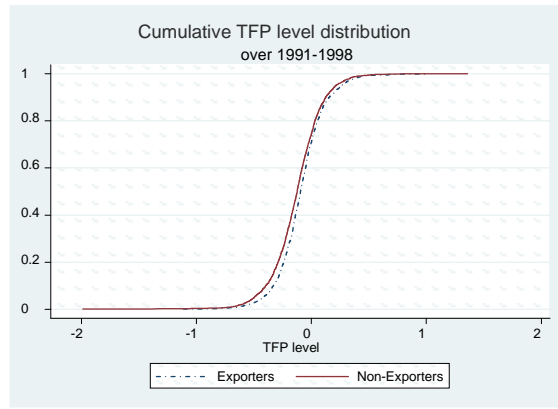
Figure 9

Cumulative Distribution of TFP Levels. Exporters Versus Non-Exporters, for Firms with R&D and without R&D, 1991-1998

R&D=1



R&D=0



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Appendix 1

Variable definitions

Variable	Description
TFP level	Firm-specific index of total factor productivity constructed using a multilateral index (the detailed explanation on its calculation is provided further in Appendix)
Export	Export status dummy, equal to 1 if firm exports at time t and 0 if it performs no exporting activities at time t
Innovation: 1) R&D investment 2) Product innovation dummy 3) Process innovation dummy 4) Innovation dummy	<ul style="list-style-type: none">- Dummy variable equal to 1 if firm invested in R&D at time $t-1$- Dummy variable equal to 1 if firm carried out product innovation only at time $t-1$; 0 – if firm performed neither product nor process innovation at time $t-1$- Dummy variable equal to 1 if firm carried out process innovation only at time $t-1$; 0 – if firm performed neither product nor process innovation at time $t-1$- Dummy variable equal to 1 if firm carried out either product or process innovation at time $t-1$; 0 – if firm performed neither product nor process innovation at time $t-1$
Foreign capital	Foreign capital dummy, equal to 1 if firm has more than 50% of foreign capital at time t
High-tech	Dummy variable, equal to 1 if firm belongs to high-tech sector

Definition of Product and Process Innovation in the ESEE Survey

Product Innovation:

- Whether a firm obtained product innovation in a given year - new products, or products with new features that are different from those that a firm produced in the previous years. If the answer is yes, the type of modification is asked:

- incorporates new materials
- incorporates new components or intermediate products
- incorporates new design or presentation
- the product performs new functions

Process Innovation:

- Whether a firm introduced an important modification in the production process. If the answer is yes, the type of modification is asked:

- introduction of new machinery
- introduction of new methods of production organization
- both

Calculation of the TFP Index

The TFP index measures the proportional difference in TFP for a firm i from size group s in year t relative to a hypothetical reference firm in the same industry r . We consider two size groups, determined by the ESEE survey. A firm belongs to a group of large firms if the number

of workers it employs is more than 200; and to a group of small firms if the number of employees is less than or equal to 200.

The reference firm is defined as follows:

- the output is equal to the geometric mean of outputs over all observations in industry r ;
- inputs are equal to the geometric means of inputs over all observations in industry r ;
- cost shares are the arithmetic means of cost shares over all observations in industry r .

The total factor productivity index for firm i ($i = 1, \dots, N$) from industry r ($r = 1, \dots, R$) and size group s in year t ($t = 1, \dots, T$) is computed using the following formula:

$$\ln TFP_{it} = \ln Y_{it} - \overline{\ln Y_{sr}} - \sum_{j=1}^J \frac{1}{2} (S_{ij} + \overline{S_{srj}}) (\ln X_{ij} - \overline{\ln X_{srj}}) + \\ + \overline{\ln Y_{sr}} - \overline{\ln Y_r} - \sum_{j=1}^J \frac{1}{2} (\overline{S_{srj}} + \overline{S_{rj}}) (\overline{\ln X_{srj}} - \overline{\ln X_{rj}})$$

where Y_{it} is an output of firm i in year t , X_{ij} is an input j ($j = \overline{1, J}$) of firm i in year t , and

S_{ij} is a cost-based share of input j of firm i in year t ; and $\overline{\ln Y_{sr}} = \frac{1}{NT} \sum_{i=1}^N \sum_{t=1}^T \ln Y_{itsr}$,

$\overline{\ln Y_r} = \frac{1}{NT} \sum_{i=1}^N \sum_{t=1}^T \ln Y_{itr}$, $\overline{\ln X_{sr}} = \frac{1}{NT} \sum_{i=1}^N \sum_{t=1}^T \ln X_{itsr}$, $\overline{\ln X_r} = \frac{1}{NT} \sum_{i=1}^N \sum_{t=1}^T \ln X_{itr}$, and

$\overline{S_{sr}} = \frac{1}{NT} \sum_{i=1}^N \sum_{t=1}^T S_{itsr}$, $\overline{S_r} = \frac{1}{NT} \sum_{i=1}^N \sum_{t=1}^T S_{itr}$ are the same variables for the reference firm.

The ESEE data provide information on the output and input variables needed to measure total factor productivity at the firm level. We model each firm as using three inputs in its production function: labor, capital and material input. The labor input is measured as the number of total effective working hours per year. The measure of capital input is the capital stock, calculated using the following formula: $k_t^* = I_t + k_{t-1}^* (1 - d_t) \frac{P_t}{P_{t-1}}$, where I represents investment in

equipment in year t , d – depreciation rates in year t , P – price indexes for equipment in year t .⁶

The material input includes raw materials, fuel and electricity costs, and other services bought by a firm. The material expenditures are deflated using the firm specific price indexes for each of the inputs provided in the ESEE survey. Firm output is defined as total firm sales corrected by inflation.

Cost-based input shares are calculated as the costs of each input in total input costs. The total input cost is the sum of the labor cost, material cost and the cost of capital. Labor costs are

⁶ The information on depreciation rates and price indexes for equipment is provided by the Instituto Nacional de Estadística (www.ine.es).

measured as total salaries to employees deflated by the consumer price index. Capital cost is computed using an estimation of the user cost of capital for each firm.

User cost of capital is calculated as the sum of the cost of long-term debt and depreciation rates less the variation of the price index for capital goods. The cost share of materials is calculated as the residual after subtracting the expenditures on labor and capital.